Fish consumption and CHD mortality: an updated meta-analysis of seventeen cohort studies

Jusheng Zheng  
Zhejiang Uni, Hangzhou, China

Tao Huang  
Zhejiang Uni, Hangzhou, China

Yinghua Yu  
University of Wollongong, yinghua@uow.edu.au

Xiaojie Hu  
Zhejiang University, China

Bin Yang

See next page for additional authors

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Abstract
Objective Results of studies on fish consumption and CHD mortality are inconsistent. The present updated meta-analysis was conducted to investigate the up-to-date pooling effects. Design A random-effects model was used to pool the risk estimates. Generalized least-squares regression and restricted cubic splines were used to assess the possible dose–response relationship. Subgroup analyses were conducted to examine the sources of heterogeneity. Setting PubMed and ISI Web of Science databases up to September 2010 were searched and secondary referencing qualified for inclusion in the study. Subjects Seventeen cohorts with 315,812 participants and average follow-up period of 15.9 years were identified. Results Compared with the lowest fish intake (<1 serving/month or 1–3 servings/month), the pooled relative risk (RR) of fish intake on CHD mortality was 0.84 (95% CI 0.75, 0.95) for low fish intake (1 serving/week), 0.79 (95% CI 0.67, 0.92) for moderate fish intake (2–4 servings/week) and 0.83 (95% CI 0.68, 1.01) for high fish intake (>5 servings/week). The dose–response analysis indicated that every 15 g/day increment of fish intake decreased the risk of CHD mortality by 6% (RR = 0.94; 95% CI 0.90, 0.98). The method of dietary assessment, gender and energy adjustment affected the results remarkably. Conclusions Our results indicate that either low (1 serving/week) or moderate fish consumption (2–4 servings/week) has a significantly beneficial effect on the prevention of CHD mortality. High fish consumption (>5 servings/week) possesses only a marginally protective effect on CHD mortality, possibly due to the limited studies included in this group.

Keywords
seventeen, fish, consumption, cohort, chd, studies, mortality, updated, meta, analysis

Disciplines
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Authors
Jusheng Zheng, Tao Huang, Yinghua Yu, Xiaojie Hu, Bin Yang, and Duo Li

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Fish consumption and CHD mortality: an updated meta-analysis of seventeen cohort studies

Jusheng Zheng¹, Tao Huang¹, Yinhua Yu¹,², Xiaojie Hu¹, Bin Yang¹ and Duo Li¹,³,∗
¹Department of Food Science and Nutrition, Zhejiang University, 268 Kaixuan Road, Hangzhou 310029, People’s Republic of China; ²School of Health Sciences, University of Wollongong, Wollongong, New South Wales, Australia; ³APCNS Centre of Nutrition and Food Safety, Hangzhou, People’s Republic of China

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Abstract

Objective: Results of studies on fish consumption and CHD mortality are inconsistent. The present updated meta-analysis was conducted to investigate the up-to-date pooling effects.

Design: A random-effects model was used to pool the risk estimates. Generalized least-squares regression and restricted cubic splines were used to assess the possible dose–response relationship. Subgroup analyses were conducted to examine the sources of heterogeneity.

Setting: PubMed and ISI Web of Science databases up to September 2010 were searched and secondary referencing qualified for inclusion in the study.

Subjects: Seventeen cohorts with 315,812 participants and average follow-up period of 15–9 years were identified.

Results: Compared with the lowest fish intake (<1 serving/month or 1–3 servings/month), the pooled relative risk (RR) of fish intake on CHD mortality was 0.84 (95 % CI 0.75, 0.95) for low fish intake (1 serving/week), 0.79 (95 % CI 0.67, 0.92) for moderate fish intake (2–4 servings/week) and 0.83 (95 % CI 0.68, 1.01) for high fish intake (>5 servings/week). The dose–response analysis indicated that every 15 g/d increment of fish intake decreased the risk of CHD mortality by 6% (RR = 0.94; 95 % CI 0.90, 0.98). The method of dietary assessment, gender and energy adjustment affected the results remarkably.

Conclusions: Our results indicate that either low (1 serving/week) or moderate fish consumption (2–4 servings/week) has a significantly beneficial effect on the prevention of CHD mortality. High fish consumption (>5 servings/week) possesses only a marginally protective effect on CHD mortality, possibly due to the limited studies included in this group.

Keywords
Fish consumption
Fatal CHD
Meta-analysis
Dose–response
Prospective cohort study

Over the past three decades there has been renewed interest in dietary intake of fish and fish oil, rich in marine-derived long-chain n-3 PUFA, and intense interest has been sparked by epidemiological studies which suggest a favourable effect of fish on CHD(1). Low mortality rate from CHD in epidemiological studies has been reported in populations with high intake of fish, such as Alaskan Natives(2), Greenland Eskimos(3) and Japanese residing in fishing villages(4). However, not all prospective cohort studies have found an inverse association between fish consumption and fatal CHD(5). In addition, clinical trials of fish or fish oil in patients with CHD did not corroborate early observational findings(6,7).

In 2004, two meta-analyses(10,11), which pooled all available cohort data at that time, reported significant inverse associations between fish consumption and fatal CHD. A consistent result was achieved in a quantitative analysis regarding this issue in 2005(12). However, stratified analyses in these meta-analyses were limited and could not help explain the heterogeneity. In addition, among several(7,13–20) recently published cohort studies, most(7,13,15,17,18,20) of them showed no statistically significant inverse association between fish intake and CHD mortality. Therefore, these results to date are still inconsistent and more detailed stratified analyses need to be done to find the potential heterogeneity. As a result, it is necessary to update the meta-analysis with relevant cohort studies in relation to fish consumption and CHD mortality.

The objective of the present study was to investigate the association between fish consumption and fatal CHD with published, up-to-date cohort studies. In addition, dose–response analysis was conducted to get trend estimation and subgroup analyses were conducted to examine sources of heterogeneity.
Methods

Selection of studies
We conducted a literature search in PubMed and ISI Web of Science up to September 2010 for all relevant papers published. Key words included ‘fishes’, ‘fish oil’, ‘seafood’, ‘omega-3’ or ‘fatty acids’ in combination with ‘coronary disease’ or ‘myocardial infarction’ (as Medical Subject Headings or text words) for the PubMed search. For the ISI Web of Science search, terms used as topic search included ‘fish’, ‘fish oil’, ‘seafood’, ‘omega-3 fatty acids’, ‘n-3 fatty acid’ or ‘polyunsaturated fatty acid’ in combination with ‘cardiovascular disease’, ‘fatal coronary heart disease’, ‘fatal myocardial infarction’ or ‘CHD’. Furthermore, references from the retrieved articles were reviewed to make sure that all the relevant bibliographies published were reviewed. Our search was confined to English-language journals only.

Two of the authors (J.Z. and T.H.) conducted the search independently. Discrepancies were resolved through group discussion. The following inclusion criteria were used in our meta-analysis: (i) studies were restricted to prospective cohort study design; (ii) risk estimates (relative risk (RR) or hazard ratio (HR)) with their corresponding 95 % confidence interval of CHD mortality rate for each category of fish consumption were provided; and (iii) if cohorts were duplicated in more than one study, the most recent and complete study (most detailed category classification) was included.

We excluded studies if: (i) cross-sectional, case–control or experimental designs were used; (ii) the outcome was not fatal CHD; (iii) there were only two fish intake categories; and (iv) the reference group was not the lowest fish intake category or the reference fish intake category was too high to be comparable with other studies.

Data extraction
We collected the following data from each publication: the first author’s name, year of publication, country of origin, range or mean of participants’ age, duration of follow-up, gender, sample size, number of events for each fish consumption category, person-years for each fish consumption category, adjusted covariates, methods of outcome assessment, methods of dietary assessment, categories of fish consumption and their corresponding RR or HR with their 95 % CI for CHD mortality. We extracted the greatest degree of adjusted risk estimates from each study.

Statistical analysis
We standardized and categorized the fish consumption into four groups based on the fish intake frequency: (i) high (>5 servings/week); (ii) moderate (2–4 servings/week); (iii) low (1 serving/week); and (iv) very low (comparison group; <1 serving/month or 1–3 servings/month). If a study contained categories of both <1 serving/month and 1–3 servings/month, we chose the previous category as the very low group. We assigned each RR from included studies into its corresponding group. If more than one fish intake category fell into the same group of our meta-analysis, we combined the RR with inverse variance weights and conducted a sensitivity analysis to examine the influence of these studies (two in low fish intake group, six in moderate fish intake group).

For the present meta-analysis, RR were used as the common risk assessment measurement and HR were considered as RR directly. RR from each study was transformed to its natural logarithm and the 95 % CI was used to calculate the corresponding standard error. Combined RR was used in a study if it presented RR with multiple outcomes or multiple exposures. The combined RR were weighted by the inverse of their variances.

All meta-analyses as well as the dose–response analysis were conducted using the STATA statistical software package version 11 (StataCorp LP, College Station, TX, USA). For dose–response analysis, we used the method described by Greenland and Longnecker and Orsini et al. to get study-specific slopes and 95 % CI from the natural logarithms of the RR and CI across categories of fish intake (nine studies). Amount of fish consumption, RR, 95 % CI and distribution of cases and person-years/non-cases in each included study should be extracted according to this method; the GLST command in STATA was used to estimate the dose–response association. If the distribution of cases and person-years/non-cases was not provided, slopes were estimated using variance-weighted least-squares regression (eight studies) and the VWLS command was used accordingly. The median or mean amount of fish consumption in each category was used for the dose–response analysis. The midpoint of upper and lower boundaries was considered as the dose of each category if the study reported only the range of fish consumption. If the highest category was open-ended, we regarded it as of the same amplitude as the preceding one. The lowest boundary was set to zero if the lowest category was open. If studies reported fish consumption using servings of fish per day, we transferred the fish amount to gram level according to the description of the study. If there was no portion size description, we deemed it to be 105 g/serving according to He et al.

To assess for potential curvilinear relation, restricted cubic splines (three knots) were used to flexibly model and graph the RR, and the MKSPLINE command, which is used for linear and restricted cubic spline construction in STATA, was chosen. To obtain a pooled RR and its 95 % CI, log RR were weighted by the inverses of their variances. A random-effects model which takes into account both within- and between-study variability was used to combine the studies. We assessed statistical heterogeneity with the $Q$ and $I^2$ statistics. $I^2$ values of 25 %, 50 % and 75 % correspond to cut-off points for low, moderate and high degrees of heterogeneity, respectively.
If heterogeneity was presented, we conducted meta-regression with study population (non-US vs. US), method of dietary assessment (interview vs. self-administered questionnaire) and energy adjustment (yes vs. no) to explore the sources of heterogeneity. Stratified analyses regarding the study region, dietary assessment, gender, follow-up period, publication year and energy adjustment were also conducted to further examine potential heterogeneity sources. A sensitivity analysis in which one study at a time was excluded was done to evaluate the effect of an individual study on the result. Funnel plots was used to assess the publication bias and Egger’s regression test to determine funnel plot asymmetry. \( P < 0.1 \) was deemed to possess publication bias. To identify and correct for funnel plot asymmetry arising from publication bias, the trim and fill algorithm was used. RR of excluded studies with only two fish intake categories (yes vs. no) or studies in which the reference group was not the lowest fish intake group (highest vs. reference) were pooled to examine the impact of these excluded studies on the overall conclusions.

Results

Search results and study characteristics

Figure 1 presents the detailed selection process. We identified twenty-nine potential studies. Seven studies were excluded because of their incomplete information on RR estimation or fish intake. Two studies were excluded because the reference group was not the lowest fish intake group (highest vs. reference) were pooled to examine the impact of these excluded studies on the overall conclusions.

Pooled effect estimates of fish consumption on fatal CHD

Pooled RR of CHD mortality in connection with low fish consumption (1 serving/week; sixteen studies included) indicated that individuals consumed 1 serving/week had...
Table 1 Characteristics of cohort studies included in the meta-analysis of fish consumption and CHD mortality

<table>
<thead>
<tr>
<th>Study source</th>
<th>Duration of follow-up (years)</th>
<th>Participants</th>
<th>Events</th>
<th>Men</th>
<th>Fish intake category</th>
<th>Exposure assessment</th>
<th>Outcome assessment</th>
<th>Adjusted variables</th>
</tr>
</thead>
<tbody>
<tr>
<td>Kromhout et al. (1985)(30), the Netherlands</td>
<td>20</td>
<td>852</td>
<td>78</td>
<td>100</td>
<td>0 g/d; 1–14 g/d; 15–29 g/d; ≥30–44 g/d; ≥45 g/d</td>
<td>Interview</td>
<td>ICD-8 (codes 410–413)</td>
<td>Age, systolic blood pressure, serum total cholesterol, cigarette smoking, subscapular skinfold thickness, physical activity, energy intake, dietary cholesterol, prescribed diet and occupation</td>
</tr>
<tr>
<td>Ascherio et al. (1995)(30), USA</td>
<td>6</td>
<td>44 895</td>
<td>264</td>
<td>100</td>
<td>&lt;1/month; 1–3/month; 1–2/week; 2–3/week; 4–5/week; ≥5/week</td>
<td>SAQ MR, AR</td>
<td></td>
<td>Age, BMI, smoking habits alcohol consumption, history of hypertension, history of diabetes, history of hypercholesterolemia, family history of myocardial infarction before 60 years of age and profession</td>
</tr>
<tr>
<td>Daviglus et al. (1997)(30), USA</td>
<td>30</td>
<td>1822</td>
<td>430</td>
<td>100</td>
<td>0 g/d; 1–17 g/d; 18–34 g/d; ≥35 g/d</td>
<td>Interview</td>
<td>ICD-8 (codes 410–414)</td>
<td>Age, education, religion, systolic pressure, serum cholesterol, no. of cigarettes smoked per day, BMI, diabetes, ECG abnormalities, daily intakes of energy, cholesterol, SFA, MUFA, PUFA, total protein, carbohydrate, alcohol, Fe, thiamin, riboflavin, niacin, vitamin C, β-carotene and retinol</td>
</tr>
<tr>
<td>Mann et al. (1997)(30), UK</td>
<td>13-3</td>
<td>10 802</td>
<td>64</td>
<td>100</td>
<td>Never*; &lt;1/week ≥1/week</td>
<td>SAQ ICD-9 (codes 410–414)</td>
<td></td>
<td>Age, sex, smoking and social class</td>
</tr>
<tr>
<td>Albert et al. (1998)(30), USA</td>
<td>11</td>
<td>20 551</td>
<td>308</td>
<td>100</td>
<td>&lt;1/month; 1–3/month; 1–2/week; 2–3/week; ≥5/week</td>
<td>Interview</td>
<td>ICD-9 (codes 410–414)</td>
<td>Age, aspirin, β-carotene treatment assignment, evidence of CVD before 12-month questionnaire, BMI, smoking status, history of diabetes, history of hypertension, history of hypercholesterolaemia, alcohol consumption, vigorous exercise, vitamin E, vitamin C and multivitamin use</td>
</tr>
<tr>
<td>Oomen et al. (2000)(50), Finland</td>
<td>20</td>
<td>1088</td>
<td>242</td>
<td>100</td>
<td>0–19 g/d; 20–39 g/d; ≥40 g/d</td>
<td>Interview</td>
<td>ICD-8 (codes 410–414, 795)</td>
<td>Age, BMI, cigarette smoking, intakes of energy, vegetables, fruit, alcohol, meat, butter and margarine</td>
</tr>
<tr>
<td>Oomen et al. (2000)(50), Italy</td>
<td>20</td>
<td>1097</td>
<td>116</td>
<td>100</td>
<td>0 g/d; 1–19 g/d; 20–39 g/d; ≥40 g/d</td>
<td>Interview</td>
<td>ICD-8 (codes 410–414, 795)</td>
<td>Age, BMI, cigarette smoking, intakes of energy, vegetables, fruit, alcohol, meat, butter and margarine</td>
</tr>
<tr>
<td>Oomen et al. (2000)(50), the Netherlands</td>
<td>20</td>
<td>553</td>
<td>105</td>
<td>100</td>
<td>0 g/d; 1–19 g/d; ≥20 g/d</td>
<td>Interview</td>
<td>ICD-8 (codes 410–414, 795)</td>
<td>Age, BMI, cigarette smoking, intakes of energy, vegetables, fruit, alcohol, meat, butter and margarine</td>
</tr>
<tr>
<td>Yuan et al. (2001)(21), China</td>
<td>12</td>
<td>18 244</td>
<td>187</td>
<td>100</td>
<td>&lt;50 g/week; ≥50 g/week</td>
<td>Interview</td>
<td>ICD-9 (codes 410–414)</td>
<td>Age, total energy intake, level of education, BMI, current smoker, average no. of cigarettes smoked per day, no. of alcoholic drinks consumed per week, history of diabetes, history of hypertension</td>
</tr>
<tr>
<td>Hu et al. (2002)(41), USA</td>
<td>16</td>
<td>84 688</td>
<td>484</td>
<td>0</td>
<td>&lt;1/month; 1–3/month; 1–2/week; 2–3/week; ≥5/week</td>
<td>SAQ MR, DC, AR</td>
<td></td>
<td>Age, time periods, smoking status, BMI, alcohol intake, menopausal status and postmenopausal hormone use, vigorous to moderate activity, no. of times aspirin was used per week, multivitamin use, vitamin E supplement use, history of hypertension, hypercholesterolaemia and diabetes, intake of trans fat, ratio of polyunsaturated fat to saturated fat and dietary fibre</td>
</tr>
<tr>
<td>Mozaffarian et al. (2003)(22), USA</td>
<td>9-3</td>
<td>3910</td>
<td>247</td>
<td>39</td>
<td>&lt;1/month; 1–3/month; 1/week; 2/week; ≥3/week</td>
<td>SAQ MR, DC</td>
<td></td>
<td>Age, gender, education, diabetes, smoking, BMI, systolic blood pressure, LDL cholesterol, HDL cholesterol, TAG, C-reactive protein, saturated fat, alcohol, meat, butter and margarine</td>
</tr>
<tr>
<td>Folsom and Demissie (2004)(13), USA</td>
<td>14</td>
<td>41836</td>
<td>922</td>
<td>0</td>
<td>&lt;0.5 g/d; 0.5–5 g/d; 5–25 g/d; ≥25 g/d</td>
<td>SAQ ICD-9 (codes 410–414,429-2) or ICD-10 (codes 120–125, IS1-6)</td>
<td></td>
<td>Age, energy intake, educational level, physical activity level, alcohol consumption, smoking status, pack-years of cigarette smoking, age at first live birth, oestrogen use, vitamin use, BMI, waist:hip ratio, diabetes, hypertension, intakes of whole grains, fruit and vegetables, red meat, cholesterol and saturated fat</td>
</tr>
<tr>
<td>Järvinen et al. (2006)(14), men, Finland</td>
<td>21-5</td>
<td>2775</td>
<td>335</td>
<td>100</td>
<td>≤11 g/d; 12–21 g/d; 22–35 g/d; 36–62 g/d; ≥63 g/d</td>
<td>Interview</td>
<td>ICD-9 (codes 410–414)</td>
<td>Age, energy intake, area, BMI, serum cholesterol, blood pressure, smoking, occupation and diabetes</td>
</tr>
<tr>
<td>Järvinen et al. (2006)(14), women, Finland</td>
<td>21-5</td>
<td>2445</td>
<td>163</td>
<td>0</td>
<td>≤8 g/d; 9–15 g/d; 16–24 g/d</td>
<td>Interview</td>
<td>ICD-9 (codes 410–414)</td>
<td>Age, gender, energy, history of hypertension and diabetes mellitus, smoking status, alcohol consumption, BMI, mental stress, walking, sports, educational level, total energy, and dietary intakes of cholesterol, saturated and n-6 polyunsaturated fatty acids, vegetables and fruit</td>
</tr>
<tr>
<td>Yamagishi et al. (2008)(13), Japan</td>
<td>12.7</td>
<td>57 972</td>
<td>419</td>
<td>39</td>
<td>0–27 g/d; 27–39 g/d; 39–53 g/d; 53–72 g/d; 72–229 g/d</td>
<td>Interview</td>
<td>ICD 10 (codes 120–125)</td>
<td>Age, energy intake, area, BMI, serum cholesterol, blood pressure, smoking, occupation and diabetes</td>
</tr>
</tbody>
</table>
Fish consumption and fatal CHD 729

Continued

Table 1

<table>
<thead>
<tr>
<th>Study source</th>
<th>Fish intake category assessment</th>
<th>Outcome assessment</th>
<th>Adjusted variables</th>
<th>Participants Events</th>
<th>Study source</th>
<th>Follow-up (years)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Goede et al. (2010)</td>
<td>&lt;3 g/d; 3–7 g/d; 7–14 g/d; 14 g/d</td>
<td>SAQ</td>
<td>Age, gender, BMI, total energy intake, ethanol intake, cigarette smoking, social economic status, vitamin or mineral supplement</td>
<td>11</td>
<td>3</td>
<td>21</td>
</tr>
<tr>
<td>Tomasello et al. (2010)</td>
<td>0–1/month; 1/month–3 g/d; 3–7 g/d; 14 g/d</td>
<td>Interview</td>
<td>Sex, age, BMI and income at study baseline</td>
<td>12</td>
<td>940</td>
<td>26</td>
</tr>
</tbody>
</table>

ASAQ, self-administered questionnaire; IC, International Classification of Diseases; MR, medical records; AR, autopsy report; DC, death certificate; ECG, electrocardiogram.

*Categories used as the very low fish consumption group (comparison group) in the meta-analysis.

Categories used as the low fish consumption group in the meta-analysis.

Categories used as the moderate fish consumption group in the meta-analysis.

Categories used as the high fish consumption group in the meta-analysis.

No heterogeneity was found ($I^2 = 0$) among studies ($P = 0.81$, 95% CI 0.70, 0.93) and exclusion of studies in which more than one category fell into the moderate fish consumption group did not markedly change the result (RR = 0.83; 95% CI 0.74, 0.94). Moderate fish consumption (2–4 servings/week; thirteen studies included) had lower CHD mortality by 21% (RR = 0.79; 95% CI 0.67, 0.92; Fig. 3) compared with the very low fish consumption; significant heterogeneity was found among studies ($P = 0.006$, $I^2 = 56.7\%$) and exclusion of studies in which more than one category fell into the moderate fish consumption group did not change the result significantly (RR = 0.80; 95% CI 0.62, 1.03). High fish consumption (>5 servings/week; five studies included) had a marginally protective effect on fatal CHD (RR = 0.83; 95% CI 0.68, 1.01; Fig. 4); no heterogeneity was found ($P = 0.451$, $I^2 = 0$).

For dose–response analysis, every 15 g/d increase of fish intake led to a significant reduction by 6% (RR = 0.94; 95% CI 0.90, 0.98) for fatal CHD (Fig. 5). However, restricted cubic splines (Fig. 6) found some evidence of a non-linear association ($P = 0.01$).

**Stratified analyses and publication bias diagnostics**

Stratified analyses were conducted to examine the sources of heterogeneity for low and moderate fish consumption, as shown in Table 3. For low and moderate fish consumption, pooled RR of studies conducted in the USA were 0.81 (95% CI 0.70, 0.93) and 0.80 (95% CI 0.70, 0.93) and consistent with the overall RR, while the heterogeneity was greatly reduced for both. The method of dietary assessment also affected the results. For moderate fish consumption, studies conducted by in-person interview shared a stronger inverse association (RR = 0.71; 95% CI 0.53, 0.96) compared with self-administered questionnaire (RR = 0.91; 95% CI 0.81, 1.02); however, the heterogeneity was greatly reduced for self-administered questionnaire ($I^2 = 23.1\%$) compared with in-person interview ($I^2 = 67.6\%$). Pooled RR estimate of fatal CHD among females was more evident than that of males for low fish consumption but not for moderate fish consumption. Pooled RR estimate of fatal CHD of studies with no energy adjustment was more evident than that with energy adjustment for both fish consumption models. For moderate model, the heterogeneity was reduced to be 0 for studies with no energy adjustment.

Pooled results of meta-regression indicated that none of the items used was a major contributor to the identified heterogeneity. Sensitivity analysis showed no significant change by excluding any study. For the low fish consumption model, visualization of the funnel plot (Fig. 7) and Egger’s test ($P = 0.265$) indicated no publication bias. For the moderate fish consumption model, visualization
**Fig. 2** (colour online) Pooled relative risk (RR) and 95% CI of studies assessing the association between low fish consumption (1 serving/week) and CHD mortality. Grey square represents the adjusted RR in each study, with the square size reflecting the study-specific weight and the 95% CI represented by horizontal bars. Open diamond indicates the pooled risk estimate and its corresponding 95% CI.

**Fig. 3** (colour online) Pooled relative risk (RR) and 95% CI of studies assessing the association between moderate fish consumption (2–4 servings/week) and CHD mortality. Grey square represents the adjusted RR in each study, with the square size reflecting the study-specific weight and the 95% CI represented by horizontal bars. Open diamond indicates the pooled risk estimate and its corresponding 95% CI.
of the funnel plot (Fig. 8) and Egger's test \((P = 0.018)\) indicated publication bias and the pooled RR remained unchanged using the trim and fill method. The pooled RR of excluded studies with only two fish intake categories was \(0.59\) (95% CI: 0.44, 0.80) and the significant inverse association still existed after adding another two excluded studies in which the reference was not the lowest fish intake category \((RR = 0.69; 95\% CI: 0.53, 0.91; \text{Fig. 9})\).
Discussion

The present study was an updated meta-analysis regarding fish consumption and CHD mortality. Its results showed a remarkable inverse association of moderate (2–4 servings/week) and low (1 serving/week) fish consumption with CHD mortality, adding updated information to the previous studies\(^{(10–12)}\). The protective effect of fish intake on CHD mortality was much stronger among those who shared the habit of moderate fish consumption (2–4 servings/week) than those who consumed low amounts of fish (1 serving/week). The protective effect of high fish consumption (>5 servings/week) on fatal CHD was much weaker and this might be due to commonly existing contaminants such as methyl mercury in the fish counterbalancing the effect of protective components. Summary RR for different categories of fish exposure did not suggest an apparent linear dose–response relationship, and a possible J-shaped relationship between fish intake and CHD mortality was indicated with restricted cubic splines. Nevertheless there were only limited high fish consumption categories, and the analysis lacked statistical power to get a definite J-shaped relationship; in addition, a linear dose–response analysis showed that every 15 g/d increment of fish consumption lowered CHD mortality by 6%, which was consistent with previous results reported by He \textit{et al.} (7% per 20 g/d increment of fish intake)\(^{(10)}\) and Konig \textit{et al.} (5.5% per 20 g/d increment of fish intake)\(^{(12)}\).

Stratified analyses

For both low and moderate fish consumption models, inverse associations between fish consumption and CHD mortality in the USA were statistically significant compared with Europe where no significant inverse associations were observed; and there was no heterogeneity in the US studies. This might be partially due to the more detailed categorization among American studies that made corresponding fish intake categories more comparable with one another, thus decreasing the heterogeneity. Moreover, the differences between the cohorts may also have contributed to the results: different populations may respond differently to dietary fish rich in \(n\)-3 PUFA based on differing genetic backgrounds. Recent emerging data documenting gene–fatty acids interactions for CHD-related traits supported this hypothesis\(^{(45)}\). Besides population regions, the method of dietary assessment also affected the pooled RR; for moderate fish consumption, self-administered questionnaire rather than in-person interview reduced the heterogeneity significantly. This is contradictory to the common understanding that in-person interview may reduce the heterogeneity more evidently than self-administered questionnaire, and the reasons are yet to be investigated. Furthermore, for both fish consumption models, summary RR with no energy adjustment were more significant than with energy adjustment. The overall protective effect of fish on fatal CHD might be decreased if all the studies are adjusted for energy. In addition, for low fish consumption, no heterogeneity was observed for studies with
Table 2: Excluded cohort studies with only two fish consumption categories or studies in which the reference group was not the lowest fish consumption category

<table>
<thead>
<tr>
<th>Study source</th>
<th>Duration of follow-up (years)</th>
<th>Participants</th>
<th>Events</th>
<th>Men</th>
<th>Fish intake category</th>
<th>RR</th>
<th>95% CI</th>
<th>Exposure assessment</th>
<th>Outcome assessment</th>
</tr>
</thead>
<tbody>
<tr>
<td>Kromhout <em>et al.</em> (1995) (36), the Netherlands</td>
<td>17</td>
<td>272</td>
<td>58</td>
<td>50</td>
<td>Yes v. no</td>
<td>Men: 0.41</td>
<td>0.20 - 0.86</td>
<td>Interview</td>
<td>ICD-9 (CODES 410–414)</td>
</tr>
<tr>
<td>Rodriguez <em>et al.</em> (1996) (38), USA (Hawaii)</td>
<td>23</td>
<td>3310</td>
<td>Unavailable</td>
<td>100</td>
<td>Yes v. no</td>
<td>Women: 0.64</td>
<td>0.25 - 1.63</td>
<td>Interview</td>
<td>MR, DC</td>
</tr>
<tr>
<td>Streppel <em>et al.</em> (2008) (17), the Netherlands</td>
<td>40</td>
<td>1373</td>
<td>348</td>
<td>100</td>
<td>Yes v. no</td>
<td>0.73</td>
<td>0.47 - 1.13</td>
<td>Interview</td>
<td>ICD-8 (codes 410–414)</td>
</tr>
<tr>
<td>Nakamura <em>et al.</em> (2005) (18), Japan</td>
<td>19</td>
<td>8879</td>
<td>142</td>
<td>80</td>
<td>&lt;1/week v. 1–2/week; 0.5/d v. 1–2/week; 1/d v. 1–2/week; &gt;2/d v. &gt;2/week</td>
<td>1.47</td>
<td>0.63 - 3.39</td>
<td>SAQ</td>
<td>ICD-9 (codes 410–414) or ICD-10 (codes 100–112)</td>
</tr>
<tr>
<td>Osler <em>et al.</em> (2003) (42), Denmark</td>
<td>11</td>
<td>7389</td>
<td>247</td>
<td>53</td>
<td>≤1/month v. 1/week; 2/month v. 1/week; ≥2/week v. 1/week</td>
<td>1.09</td>
<td>0.78 - 1.52</td>
<td>SAQ</td>
<td>ICD-8 (codes 410–414)</td>
</tr>
</tbody>
</table>

RR, relative risk; SAQ, self-administered questionnaire; ICD, International Classification of Diseases; MR, medical records; DC, death certificate.

Advantages and Limitations

The present meta-analysis possesses several advantages over previous works. First, it included more detailed stratified analyses, allowing for a more sophisticated examination of potential heterogeneity. Second, it was able to pool RR based on more detailed fish consumption categories and portion sizes, providing more accurate estimates than previous works. Third, it addressed residual confounding by adjusting for more variables. Fourth, it conducted a more comprehensive data synthesis, including studies with different designs and outcomes. Fifth, it provided a more comprehensive meta-analysis with greater statistical power. However, some limitations should be noted. First, the limited studies among high fish consumption categories could not be categorized into the high fish consumption group (five studies) owing to the fact that many studies assessed only limited fish intake categories, which could not be categorized into the high fish consumption category. Second, the high fish consumption category might have been overestimated by the inclusion of studies with only two fish consumption categories or studies in which the reference group was not the lowest fish consumption category. Third, the limited studies about women might be partially due to the limited studies among high fish consumption categories and gives a clear idea about the impact of fish consumption categories among high fish consumption categories.

In conclusion, the present meta-analysis provides a more comprehensive understanding of the relationship between fish consumption and fatal CHD. Future studies should focus on more detailed fish consumption categories and portion sizes, and should include more studies about women. More studies regarding the effects of fish consumption on fatal CHD in women are needed to clarify the gender differences.
weight. This procedure might lead to overestimation of the precision of the RR estimates; however, for both low and moderate fish consumption groups, our sensitivity analyses did not show any significant changes excluding these studies. In addition, for moderate fish consumption (2–4 servings/week), publication bias (trim and fill method did not change the pooled RR) was found in the meta-analysis and great heterogeneity was also observed; these might be partially due to the relatively wide range of moderate fish consumption (2–4 servings/week). We included papers published only in the English language, which would be the source of our publication bias.

### Table 3

<table>
<thead>
<tr>
<th>Subgroups</th>
<th>Low fish consumption (1 serving/week)</th>
<th>Moderate fish consumption (2–4 servings/week)</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>n</td>
<td>RR</td>
</tr>
<tr>
<td>All studies</td>
<td>16</td>
<td>0.84</td>
</tr>
<tr>
<td>Region</td>
<td></td>
<td></td>
</tr>
<tr>
<td>USA</td>
<td>7</td>
<td>0.81</td>
</tr>
<tr>
<td>Europe</td>
<td>8</td>
<td>0.95</td>
</tr>
<tr>
<td>Finland</td>
<td>3</td>
<td>0.99</td>
</tr>
<tr>
<td>Netherlands</td>
<td>3</td>
<td>0.76</td>
</tr>
<tr>
<td>Asia</td>
<td>1</td>
<td>0.67</td>
</tr>
<tr>
<td>Follow-up period (years)</td>
<td></td>
<td></td>
</tr>
<tr>
<td>≤10</td>
<td>2</td>
<td>0.94</td>
</tr>
<tr>
<td>10–20</td>
<td>7</td>
<td>0.74</td>
</tr>
<tr>
<td>≥20</td>
<td>7</td>
<td>0.94</td>
</tr>
<tr>
<td>Dietary assessment</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Interview</td>
<td>9</td>
<td>0.88</td>
</tr>
<tr>
<td>Self-administered FFQ</td>
<td>7</td>
<td>0.82</td>
</tr>
<tr>
<td>Gender</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Male</td>
<td>9</td>
<td>0.90</td>
</tr>
<tr>
<td>Female</td>
<td>3</td>
<td>0.79</td>
</tr>
<tr>
<td>Both</td>
<td>4</td>
<td>0.75</td>
</tr>
<tr>
<td>Published year</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Before 2004</td>
<td>12</td>
<td>0.83</td>
</tr>
<tr>
<td>After 2004</td>
<td>5</td>
<td>0.80</td>
</tr>
<tr>
<td>Energy adjustment</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Yes</td>
<td>10</td>
<td>0.87</td>
</tr>
<tr>
<td>No</td>
<td>6</td>
<td>0.79</td>
</tr>
</tbody>
</table>

**RR**, relative risk.

![Begg's funnel plot](image)
Apart from the above issues, the cooking methods and contaminants (e.g., methyl mercury) of certain types of fish affect the results severely. However, among included studies, few\(^5,22\) contained information about these issues.

**Mechanisms of fish on fatal CHD**

It has long been assumed that long-chain \(n\)-3 PUFA, including \(20:5n-3\) and \(22:6n-3\), play an important role in the protective effect of fish on CHD risk. The possible mechanisms include their antiarrhythmic properties, reduction of serum TAG\(^{46}\) and platelet aggregation\(^{47}\). Fish oil may also improve endothelial dysfunction\(^{48}\), which is considered an early marker of atherosclerosis. However, considering the synergic effect of many components in fish, such as high-quality protein, amino acid and vitamins, analysis of total fish consumption on CHD is probably

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**Fig. 8** (colour online) Begg’s funnel plot with pseudo 95% confidence limits indicating the publication bias of the relative risk (RR) assessing the association of moderate fish consumption (2–4 servings/week) and CHD mortality. The horizontal line indicates the summary estimate of RR, with the sloping dashed lines representing the expected 95% CI for a given SE

![Funnel plot](image)

**Study** | **RR (95% CI)**
---|---
Study with only two categories (yes v. no) | 0.41 (0.20, 0.85)
Kromhout et al. (1995)\(^{36}\), men, the Netherlands | 0.64 (0.25, 1.63)
Kromhout et al. (1995)\(^{36}\), women, the Netherlands | 0.50 (0.28, 0.90)
Rodriguez et al. (1996)\(^{39}\), USA (Hawaii) | 0.73 (0.47, 1.13)
Streppel et al. (2008)\(^{17}\), the Netherlands | 0.59 (0.44, 0.80)
Subtotal \((I^2 = 0.0\%, P = 0.533)\) | 0.98 (0.63, 1.53)

**Study in which the reference group was not the lowest one (highest v. reference)**

<table>
<thead>
<tr>
<th>Study</th>
<th>RR (95% CI)</th>
</tr>
</thead>
</table>
Osler et al. (2003)\(^{42}\), Denmark | 0.91 (0.35, 2.36)
Nakamura et al. (2005)\(^{18}\), Japan | 0.97 (0.64, 1.45)
Subtotal \((I^2 = 0.0\%, P = 0.890)\) | 0.98 (0.63, 1.53)
Overall \((I^2 = 14.4\%, P = 0.322)\) | 0.69 (0.53, 0.91)

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**Fig. 9** (colour online) Pooled relative risk (RR) and 95% CI of excluded studies with only two fish intake categories (yes v. no) or studies in which the reference group was not the lowest fish intake category (highest v. reference) in assessing the association between fish consumption and CHD mortality. Grey square represents the adjusted RR in each study, with the square size reflecting the study-specific weight and the 95% CI represented by horizontal bars. Open diamond indicates the pooled risk estimate and its corresponding 95% CI.
more valuable than the sole evaluation of long-chain n-3 PUFA\(^{(49)}\).

**Conclusion**

In conclusion, fish consumption of 1 serving/week or 2–4 servings/week has a significant protective effect on fatal CHD; fish consumption of > 5 servings/week could marginally decrease CHD mortality, but the limited number of studies included in this group might contribute to the result. Our findings support the public dietary guideline to eat fish two times per week.

**Acknowledgements**

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**References**

Fish consumption and fatal CHD